



Pearson
Edexcel

Mark Scheme (Results)

Autumn 2020

Pearson Edexcel GCE In A Level Statistics
(9ST0/02)

Paper 2: Statistical Inference

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Autumn 2020

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General Marking Guidance

Total marks

The total number of marks for the paper is 80.

Mark types

The Edexcel Statistics mark schemes use the following types of marks:

- **M** **Method** marks, awarded for ‘knowing a method and attempting to apply it’, unless otherwise indicated.
- **A** **Accuracy** marks can only be awarded if the relevant method (M) marks have been earned.
- **B** **Unconditional accuracy** marks are independent of M marks
- **E** **Explanation** marks

NOTE: Marks should not be subdivided.

Abbreviations

These are some of the marking abbreviations that will appear in the mark schemes.

- ft follow through
- PI possibly implied
- cao correct answer only
- cso correct solution only
(There must be no errors in this part of the question)
- awrt answers which round to
- awfw answers which fall within (a given range)
- SC special case
- nms no method shown
- oe or equivalent
- dep dependent (on a given mark or objective)
- dp decimal places
- sf significant figures
- * The answer is printed on the paper

Further notes

- All candidates must receive the same treatment. Examiners must mark the first candidate in exactly the same way as they mark the last.
- Mark schemes should be applied **positively**. Candidates must be rewarded for what they have shown they can do rather than penalised for omissions.
- Examiners should mark according to the mark scheme not according to their perception of where the grade boundaries may lie.
- There is **no ceiling** on achievement. All marks on the mark scheme should be used appropriately.
- All the marks on the mark scheme are designed to be awarded. Examiners should always award full marks if deserved, i.e. if the answer matches the mark scheme. Examiners should also be prepared to award zero marks if the candidate's response is not worthy of credit according to the mark scheme.
- Where some judgement is required, mark schemes will provide the principles by which marks will be awarded and exemplification may be limited.
- When examiners are in doubt regarding the application of the mark scheme to a candidate's response, the team leader must be consulted.
- All A marks are 'correct answer only' (cao.), unless shown, for example, as A1ft to indicate that previous wrong working is to be followed through.
- After a **misread**, the subsequent A marks affected are treated as A1ft, but manifestly absurd answers should never be awarded A marks.
- **Crossed out** work should be marked UNLESS the candidate has replaced it with an alternative response.
- If **two solutions** are given, each should be marked, and the resultant mark should be the mean of the two marks, rounded down to the nearest integer if needed.

| Question | Scheme | Marks | AO | Notes |
|--------------|---|-----------|---------------|------------------------------------|
| 1(a) | $6.3 \pm z \times \frac{0.74}{\sqrt{14}}$ | M1 | 1.3 | PI (z or t) |
| | $z = 1.96$ (or 2.0 to 2 s.f.) | B1 | 1.3 | PI or $t=2.160$ |
| | CI is (5.91, 6.69) | A1 | 1.3 | awrt (5.9, 6.7) |
| 1(b) | 5.8 is not in the confidence interval | B1ft | 2.1b | oe comparison of 5.8 with their CI |
| | There is significant evidence that the mean weight of baby girls at 3 months of age is different from 5.8 kg | E1dep | 2.1a | oe Dep on previous B1 |
| 1(c) | It is a random sample of babies weights at 3 months of age in Rose's clinic | | | oe |
| | The babies' weights at 3 months of age in Rose's clinic are normally distributed | | | oe |
| | The standard deviation of the babies' weights at 3 months of age in Rose's clinic is the same as the standard deviation for UK babies | | | oe |
| | | E1, E1 | 3.1a, 3.1a | Any 2 for 2 marks |
| Total | | 7 | | |

| Question | Scheme | Marks | AO | Notes |
|----------|--|-------|------|------------------------------------|
| 2 | Exact binomial method | | | |
| | $H_0: \pi = 0.75$ $H_1: \pi > 0.75$ | B1 | 1.3 | Both correct Condone use of 'p' |
| | $n = 45$ so let $X \sim B(45, 0.75)$ | B1 | 1.3 | Use of Binomial with $n=45$ |
| | $P(X \geq 36)$ | M1 | 1.3 | |
| | $= 1 - P(X \leq 35)$ | M1 | 1.3 | |
| | $= 0.27997\dots$ | A1 | 1.3 | AWRT 0.28 or CR: $X \geq 39$ |
| | $(0.28 > 0.05)$ so do not reject H_0 | A1dep | 2.1b | PI Dep on correct ts/ cv oe |
| | Julia's survey result does not provide evidence that the percentage preferring unwrapped cucumbers is higher than 75% or There is insufficient evidence members of her gardening club have a greater preference for unwrapped cucumbers. | E1dep | 2.1a | Dep on correct ts/ cv oe |
| | | | | |

| Question | Scheme | Marks | AO | Notes |
|----------|--|---------|----|--|
| | Critical Region method | | | |
| | $H_0: \pi = 0.75$ $H_1: \pi > 0.75$ | (B1) | | Both correct Condone use of 'p' |
| | $n = 45$ so let $X \sim B(45, 0.75)$ | (B1) | | Use of Binomial with $n=45$ |
| | $P(X \geq 39) = 0.044$ | (M1) | | Attempt to evaluate $P(X \geq x)$ to find CR |
| | | (M1) | | $P(X \geq 39)$ calculated condone $P(X \geq 38) = 0.094$ |
| | CR: $X \geq 39$ | (A1) | | CR: $X \geq 39$ |
| | $(36 < 39)$ so do not reject H_0 | (A1dep) | | PI Dep on correct ts/ cv oe |
| | Julia's survey result does not provide evidence that the percentage preferring unwrapped cucumbers is higher than 75% or There is insufficient evidence members of her gardening club have a greater preference for unwrapped cucumbers. | (E1dep) | | Dep on correct ts/ cv oe |
| | | | | |

| Question | Scheme | Marks | AO | Notes |
|-------------|--|----------|----|---|
| 2 (cont) | Normal approximation method | | | |
| | $H_0: \pi = 0.75$ $H_1: \pi > 0.75$ | (B1) | | Both correct Condone use of 'p' |
| | $\hat{p} = \frac{36}{45} = 0.8$ | (B1) | | |
| | $test\ statistic = \frac{0.8 - 0.75}{\sqrt{\frac{0.75 \times 0.25}{45}}}$ | (M1) | | PI Correct numerator Allow actual numbers rather than proportions |
| | | (M1) | | Correct denominator |
| | $test\ statistic = 0.775$ or $p\text{-value} = 0.219\dots$ | (A1) | | ts AFWW 0.77 - 0.78 or $p\text{-value}$ |
| | $(0.775 < 1.6649)$ or $0.219\dots > 0.05$ so do not reject H_0 | (A1dep) | | PI Dep on correct ts/ cv oe |
| | Julia's survey result does not provide evidence that the percentage preferring unwrapped cucumbers is higher than 75% or There is insufficient evidence members of her gardening club have a greater preference for unwrapped cucumbers. | (E1dep) | | Dep on correct ts/ cv oe |
| | Total | 7 | | |

| Question | Scheme | Marks | AO | Notes |
|----------|--|-------|------|--|
| 3(a) | $\alpha = 0.05$ OR $df = n - 1 = 24$ | B1 | 1.3 | Mark for use of $\alpha=0.05$ if p -value method used or for correct df if critical value method used PI |
| | p -value = 0.00752... used with 0.05 or $t = 2.619$ used with $t_{24}(0.05) = 1.71$ | M1 | 1.3 | |
| | $0.00752 < 0.05$ or $2.619 \dots > t_{24}(0.05) = 1.71$ \therefore reject H_0 | A1dep | 2.1b | Dep B1 M1 |
| | There is evidence at the 5% level that the children did better , on average, in the test after practising the skill. | E1dep | 2.1a | cso Dep on all previous marks |
| 3(b) | Type I error is rejecting H_0 when H_0 is true | E1 | 3.1a | Possibly stated in context |
| | $P(\text{type I error}) = 0.05$ | A1 | 2.1a | |

| Question | Scheme | Marks | AO | Notes |
|-----------------|--|--------------|-----------|---------------------|
| 3(c) | The conclusion may not be reliable as... | E1 | 3.1b | Any sensible reason |
| | The differences in scores may not be normally distributed | | | |
| | The improvement in scores may not have been due to the computer program. | | | |
| | The pupils may not be representative of the population as a whole. | | | |
| | The conclusion is likely to be reliable as... | | | |
| | The 25 children are a random sample (of program users) | | | |
| Total | | 6 | | |

| Question | Scheme | | | Marks | AO | Notes | | | | | | | | | | | | | | | |
|----------------------------|---|-------|---------------------|-------|---|---|----|----|--------------|---|---|----------------|----|----|----------------------------|---|---|--|--|--|--|
| 4(a) | $\frac{5 \times 57}{95} = 3$ | 2 | | M1 | 1.3 | Attempted method for calculating at least one expected value - PI | | | | | | | | | | | | | | | |
| | 18 | 12 | | | | | | | | | | | | | | | | | | | |
| | 3 | 2 | | | | | | | | | | | | | | | | | | | |
| | 1.8 | 1.2 | | | | | | | | | | | | | | | | | | | |
| | 7.8 | 5.2 | | | | | | | | | | | | | | | | | | | |
| | 23.4 | 15.6 | | | | | | | | | | | | | | | | | | | |
| | 57 | 38 | | | | | | | | | | | | | | | | | | | |
| | Expected values < 5 for Australia & South Africa, Far East and South America So these regions need to be combined with each other (or others) | | | E1 | 3.1a | | | | | | | | | | | | | | | | |
| 4(b) | H_0 : no association (between region of the world and sex of player) H_1 : an association (between region of the world and sex of player) | | | B1 | 1.3 | Both correct | | | | | | | | | | | | | | | |
| | Observed frequencies: | | | | | | | | | | | | | | | | | | | | |
| | <table border="1"> <thead> <tr> <th>Region of the World</th> <th>Men</th> <th>Women</th> </tr> </thead> <tbody> <tr> <td>Eastern Europe</td> <td>11</td> <td>19</td> </tr> <tr> <td>USA & Canada</td> <td>8</td> <td>5</td> </tr> <tr> <td>Western Europe</td> <td>29</td> <td>10</td> </tr> <tr> <td>Other Regions of the World</td> <td>9</td> <td>4</td> </tr> </tbody> </table> | | Region of the World | Men | Women | Eastern Europe | 11 | 19 | USA & Canada | 8 | 5 | Western Europe | 29 | 10 | Other Regions of the World | 9 | 4 | | | | |
| Region of the World | Men | Women | | | | | | | | | | | | | | | | | | | |
| Eastern Europe | 11 | 19 | | | | | | | | | | | | | | | | | | | |
| USA & Canada | 8 | 5 | | | | | | | | | | | | | | | | | | | |
| Western Europe | 29 | 10 | | | | | | | | | | | | | | | | | | | |
| Other Regions of the World | 9 | 4 | | | | | | | | | | | | | | | | | | | |
| | | | M1 | 1.3 | Observed frequencies required for combined (Other regions) class only PI | | | | | | | | | | | | | | | | |

| Question | Scheme | Marks | AO | Notes | | | | | | | | | | | | | | | |
|----------------------------|---|---------------------|--|------------------------|----------------|-------|-------|--------------|-------|-------|----------------|-------|-------|----------------------------|-------|-------|--|--|--|
| 4(b) (cont) | Expected frequencies: <table border="1"> <thead> <tr> <th>Region of the World</th> <th>Men</th> <th>Women</th> </tr> </thead> <tbody> <tr> <td>Eastern Europe</td> <td>18</td> <td>12</td> </tr> <tr> <td>USA & Canada</td> <td>7.8</td> <td>5.2</td> </tr> <tr> <td>Western Europe</td> <td>23.4</td> <td>15.6</td> </tr> <tr> <td>Other Regions of the World</td> <td>7.8</td> <td>5.2</td> </tr> </tbody> </table> | Region of the World | Men | Women | Eastern Europe | 18 | 12 | USA & Canada | 7.8 | 5.2 | Western Europe | 23.4 | 15.6 | Other Regions of the World | 7.8 | 5.2 | | | |
| Region of the World | Men | Women | | | | | | | | | | | | | | | | | |
| Eastern Europe | 18 | 12 | | | | | | | | | | | | | | | | | |
| USA & Canada | 7.8 | 5.2 | | | | | | | | | | | | | | | | | |
| Western Europe | 23.4 | 15.6 | | | | | | | | | | | | | | | | | |
| Other Regions of the World | 7.8 | 5.2 | | | | | | | | | | | | | | | | | |
| | A1 | 1.3 | Expected values for all classes but may appear in 4(a) above PI awrt values in table | | | | | | | | | | | | | | | | |
| | Contribution to χ^2: <table border="1"> <thead> <tr> <th>Region of the World</th> <th>Men</th> <th>Women</th> </tr> </thead> <tbody> <tr> <td>Eastern Europe</td> <td>2.722</td> <td>4.083</td> </tr> <tr> <td>USA & Canada</td> <td>0.005</td> <td>0.007</td> </tr> <tr> <td>Western Europe</td> <td>1.340</td> <td>2.010</td> </tr> <tr> <td>Other Regions of the World</td> <td>0.185</td> <td>0.277</td> </tr> </tbody> </table> | Region of the World | Men | Women | Eastern Europe | 2.722 | 4.083 | USA & Canada | 0.005 | 0.007 | Western Europe | 1.340 | 2.010 | Other Regions of the World | 0.185 | 0.277 | | | |
| Region of the World | Men | Women | | | | | | | | | | | | | | | | | |
| Eastern Europe | 2.722 | 4.083 | | | | | | | | | | | | | | | | | |
| USA & Canada | 0.005 | 0.007 | | | | | | | | | | | | | | | | | |
| Western Europe | 1.340 | 2.010 | | | | | | | | | | | | | | | | | |
| Other Regions of the World | 0.185 | 0.277 | | | | | | | | | | | | | | | | | |
| | M1 | 1.3 | Attempt at $\frac{(O-E)^2}{E}$ PI | | | | | | | | | | | | | | | | |
| | Test stat = $\frac{(11-18)^2}{18} + \dots + \frac{(4-5.2)^2}{5.2}$ | M1 | 1.3 | Intention to sum PI | | | | | | | | | | | | | | | |
| | $\chi^2 = 10.63$ | A1 | 1.3 | awrt 10.6 | | | | | | | | | | | | | | | |

| Question | Scheme | Marks | AO | Notes |
|------------------------|--|--------------|-----------|--|
| 4(b) (cont) | $p\text{-value} = 0.0139$ or cv of χ^2 at 5% level = 7.81 | M1 | 1.3 | $p\text{-value}$ must be compared with $\alpha = 0.05$ |
| | $10.63 > 7.81$ or $0.0139 < 0.05$ so reject H_0 | A1dep | 2.1b | Dep on ts/ cv or $p\text{-value}$ correct PI |
| | There is significant evidence of an association between region of the world and sex of player | E1dep | 2.1a | Dep on ts/ cv or $p\text{-value}$ correct |
| 4(c) | The greatest contribution (4.08) to the association is from women and 'Eastern Europe' where more women were observed (19) than would be expected (12) from Eastern Europe to earn at least \$1 million in 2018. | E1ft | 2.1a | Or fewer men from Eastern Europe No numerical justification |
| | | E1ft | 2.1b | Full numerical justification |
| Total | | 14 | | |

| Question | Scheme | Marks | AO | Notes |
|--------------|--|-----------|------|--|
| 5(a) | Sign Test | B1 | 2.1a | |
| | $H_0: \text{population median} = -0.02$ $H_1: \text{population median} > -0.02$ | B1 | 1.3 | Condone use of η or two-tail H_1 May be awarded if seen in (b) |
| | $P(x \leq 2) = 0.0547 > 0.05$ Do not reject H_0 | E1 | 2.1b | |
| 5(b) | $H_0: \text{population average} = -0.02$ $H_1: \text{population average} > -0.02$ | B1 | 1.3 | One-tail H_1 used Allow mean or median oe μ/η |
| | Ranks of $ X $: 2 1 3 4 5 6 7 8 9 10 | M1 | 1.3 | Only first two required check table |
| | $W = 1 + 2 = 3$ | A1 | 1.3 | Or 52 |
| | $cv = 11$ (for $\alpha = 0.05$, one-tailed) | B1 | 1.3 | Or 44 |
| | $3 < 11$ so reject H_0 (one-tailed) | M1 | 2.1b | Same tail comparison |
| | Sufficient evidence that the average performance for SW London is better than the average for England. | E1 | 2.1a | Allow mean or median Dep on correct ts/cv |
| 5(c) | If the ranks are consistent with a symmetric population then Wilcoxon can be used. (Otherwise the sign test is needed). | E1 | 3.1a | |
| | Use Wilcoxon if possible because it is a more powerful test OR Use Wilcoxon if possible as it takes account of the sizes of the differences rather than just the +/- signs | E1 | 3.1a | |
| Total | | 11 | | |

| Question | Scheme | Marks | AO | Notes |
|----------|--|-------|------|---|
| 6(a) | For a double-blind trial neither the patients nor the doctors/researchers should know which treatment has been assigned to any patient. | E1 | 1.1 | |
| | Patients would be aware of which diet/treatment they were receiving so it couldn't be a double-blind trial. | E1 | 1.1 | |
| 6(b) | $H_0: \mu_A = \mu_B$ $H_1: \mu_A \neq \mu_B$ | B1 | 1.3 | both |
| | $\text{Test stat} = \frac{10.7 - 3.1}{\sqrt{\left(\frac{9.6^2}{104} + \frac{7.0^2}{95}\right)}} = 6.419..$ | M1 | 1.3 | $10.7 - 3.1$ |
| | | M1 | 1.3 | $\sqrt{\left(\frac{9.6^2}{104} + \frac{7.0^2}{95}\right)}$ |
| | Test stat = 6.419.. | A1 | 1.3 | AWRT 6.42 Ignore sign |
| | Critical value = ± 1.96 OR p -value $P(z > 6.419..) = 1.37 \times 10^{-10}$ | B1 | 1.3 | $Z > 6.419$ implied by correct p -value |
| | $(6.419 > 1.96 \text{ or } 1.37 \times 10^{-10} < 0.025)$ so reject H_0 | A1dep | 2.1b | cv correct and compared with ts OR p -value compared to 0.025 PI Dep on cv/ ts or p -value correct |
| | There is significant evidence of a difference in mean weight loss between patients assigned to Diet A and those assigned to Diet B | E1dep | 2.1a | Conclusion correct and in context; test all correct Dep on cv/ ts or p -value correct |
| | | | | |
| | | | | |

| Question | Scheme | Marks | AO | Notes |
|----------|--|---------|------|---|
| | Alternative | | | |
| | $H_0: \mu_A = \mu_B$ $H_1: \mu_A \neq \mu_B$ | (B1) | | oe both |
| | Test stat = $\frac{10.7 - 3.1}{\sqrt{\left(\frac{8.46^2}{104} + \frac{8.46^2}{95}\right)}}$ = 6.33.. | (M1) | | 10.7 – 3.1 |
| | $\sqrt{\left(\frac{8.46^2}{104} + \frac{8.46^2}{95}\right)}$ | (M1) | | Use their s_p^2 |
| | Test stat = 6.33.. | (A1) | | AWRT 6.33 Ignore sign |
| | Critical value = ± 1.972 OR p -value $P(z > 6.33..) = 1.62 \times 10^{-9}$ | (B1) | | $z > 6.33$ implied by correct p -value |
| | (6.33 > 1.97 or $1.62 \times 10^{-9} < 0.025$) so reject H_0 | (A1dep) | | cv correct and compared with ts OR p -value compared to 0.025 PI Dep on cv/ ts or p -value correct |
| | There is significant evidence of a difference in mean weight loss between patients assigned to Diet A and those assigned to Diet B | (E1dep) | | Conclusion correct and in context; test all correct Dep on cv/ ts or p -value correct |
| 6(c) | The sample means are approximately normally distributed large samples so CLT applies | E1 | 3.1a | Large samples, CLT |

| Question | Scheme | Marks | AO | Notes |
|----------|---|-------|------|---|
| 6(d) | Patients did not deviate from diet | | | |
| | Same scales used | | | |
| | Sample variances can be used in place of unknown population variances (large samples) | | | |
| | | E1 | 3.1a | Any sensible explanation |
| 6(e) | $H_0: \mu_A - \mu_B = 4$ $H_1: \mu_A - \mu_B > 4$ | B1 | 1.3 | oe Both |
| | $\text{Test stat} = \frac{10.7 - 3.1 - (4)}{\sqrt{\left(\frac{9.6^2}{104} + \frac{7.0^2}{95}\right)}} = 3.040..$ | M1 | 1.3 | $7.6 - (4 - 0)$ |
| | Test stat = 3.04 | A1 | 1.3 | AWRT 3.04 |
| | Critical value = 1.6449 OR p-value $P(z > 3.04) = 0.00118$ | B1 | 1.3 | $Z > 3.04$ implied by correct p-value |
| | $(3.04 > 1.6449 \text{ or } 0.00118 < 0.05)$ so reject H_0 | A1dep | 2.1b | PI Dep on cv/ ts or p-value correct |
| | There is significant evidence that the mean weight loss of patients assigned to Diet A is at least 4kg more than that of patients assigned to Diet B. OR There is significant evidence the mean weight loss of patients assigned to Diet A is medically worthwhile | E1dep | 2.1a | Conclusion correct and in context; test all correct Dep on cv/ ts or p-value correct |

| Question | Scheme | Marks | AO | Notes |
|----------|---|-----------|----|--|
| 6(e) | Alternative | | | |
| | $H_0: \mu_A - \mu_B = 4$ $H_1: \mu_A - \mu_B > 4$ | (B1) | | oe Both |
| | Test stat = $\frac{10.7 - 3.1 - (4)}{\sqrt{\left(\frac{8.46^2}{104} + \frac{8.46^2}{95}\right)}} = 2.99..$ | (M1) | | $7.6 - (4 - 0)$ |
| | Test stat = 2.99... | (A1) | | AWRT 3.0 |
| | Critical value = 1.653 OR p -value ($z > 2.99$) = 0.001395 | (B1) | | $z > 2.99$ implied by correct p -value |
| | ($2.99 > 1.653$ or $0.001395 < 0.05$) so reject H_0 | (A1dep) | | PI Dep on cv/ ts or p -value correct |
| | There is significant evidence that the mean weight loss of patients assigned to Diet A is at least 4kg more than that of patients assigned to Diet B. OR There is significant evidence the mean weight loss of patients assigned to Diet A is medically worthwhile | (E1dep) | | Conclusion correct and in context; test all correct Dep on cv/ ts or p -value correct |
| | Total | 17 | | |

| Question | Scheme | Marks | AO | Notes | | | | | | | | | | | | | | | | | | | | |
|---------------------|---|---------------------|--------|---------------------------------------|-----|-------|----------------|---|---------|--------|-------|-------|----|---------|--------|--|-------|----|---------|--|--|--|--|--|
| 7(a) | $H_0: \mu_A = \mu_B = \mu_C = \mu_D$ $H_1: \text{at least two of } \mu_A, \mu_B, \mu_C, \mu_D \text{ are different}$ | B1 | 2.1a | or 1-factor ANOVA stated | | | | | | | | | | | | | | | | | | | | |
| | $T = 654.528$ $\sum \sum x_{ij}^2 = 15508.274$ | B1 | 1.3 | Either | | | | | | | | | | | | | | | | | | | | |
| | $SS_T = \sum \sum x_{ij}^2 - \frac{T^2}{n}$ $= 15508.274 - \frac{654.528^2}{29}$ $= 735.622$ | M1 | 1.3 | SS Total | | | | | | | | | | | | | | | | | | | | |
| | $SS_B = \sum \frac{T_i^2}{n_i} - \frac{T^2}{n}$ $SS_B = \frac{190.078^2}{7} + \frac{184.401^2}{8}$ $+ \frac{133.191^2}{6}$ $+ \frac{146.858^2}{8}$ $- \frac{654.528^2}{29}$ $= 291.742$ | M1 | 1.3 | SS between tyre brands | | | | | | | | | | | | | | | | | | | | |
| | $SS_E = 735.622 - 291.742 = 443.880$ | M1ft | 1.3 | | | | | | | | | | | | | | | | | | | | | |
| | <table border="1"> <thead> <tr> <th>Source of variation</th> <th>df</th> <th>SS</th> <th>MSS</th> <th>Ratio</th> </tr> </thead> <tbody> <tr> <td>Between brands</td> <td>3</td> <td>291.742</td> <td>97.247</td> <td>5.477</td> </tr> <tr> <td>Error</td> <td>25</td> <td>443.879</td> <td>17.755</td> <td></td> </tr> <tr> <td>Total</td> <td>28</td> <td>735.621</td> <td></td> <td></td> </tr> </tbody> </table> | Source of variation | df | SS | MSS | Ratio | Between brands | 3 | 291.742 | 97.247 | 5.477 | Error | 25 | 443.879 | 17.755 | | Total | 28 | 735.621 | | | | | |
| Source of variation | df | SS | MSS | Ratio | | | | | | | | | | | | | | | | | | | | |
| Between brands | 3 | 291.742 | 97.247 | 5.477 | | | | | | | | | | | | | | | | | | | | |
| Error | 25 | 443.879 | 17.755 | | | | | | | | | | | | | | | | | | | | | |
| Total | 28 | 735.621 | | | | | | | | | | | | | | | | | | | | | | |
| | | B1 | 1.3 | df correct | | | | | | | | | | | | | | | | | | | | |
| | | M1 | 1.3 | MS=SS/df for between brands and error | | | | | | | | | | | | | | | | | | | | |
| | $F = \frac{97.247}{17.755} = 5.477$ | A1 | 1.3 | AWRT 5.5 OR p = 0.0049 | | | | | | | | | | | | | | | | | | | | |

| Question | Scheme | Marks | AO | Notes |
|----------------|---|-------|------|--|
| 7(a) (cont) | cv: $F_{25}^3(0.05) = 2.991$ or $F_{25}^3(0.01) = 4.675$ | B1 | 1.3 | Either cv |
| | 5.477 > cv so reject H_0 | M1 | 2.1b | or $p=0.0049<0.05$ or 0.01 PLUS correct conclusion |
| | There is significant evidence to suggest that at least two of the four mean lives of the tyre brands are different. | A1 | 2.1a | Conclusion in context condone statement that brands A and D differ |

| Question | Scheme | Marks | AO | Notes |
|-----------------|---|--------------|-----------|---|
| 7(b) | Possible comments | | | |
| | Scatter diagram | | | |
| | The scatter diagram is consistent with normality within tyre brands so there is no reason to doubt the validity of the test on this basis. | | | |
| | The scatter diagram appears to show similar spread (variances) for all tyre brands so there is no reason to doubt the validity of the test on this basis. | | | Condone brand B has a greater spread so not valid |
| | | E1 | 3.1a | Either of these |
| | Data collection | | | |
| | The tyre sample should be a completely randomised design (CRD). There may be bias due to the opportunistic nature of Daniel's sampling method. | | | |
| | Larger samples would have enabled tests for normality and comparisons of variances within 'treatments'. | | | |
| | | E1 | 3.1a | Either of these |
| | | E1 | 3.1a | One extra comment from either category |

| Question | Scheme | Marks | AO | Notes | | | | | | | | | | | | | | | | | | | | |
|--------------|---|-----------|-------|---|--|--|--|---|---|---|---|--------------|--|--|--|--|-------------|--|--|--|--|----|------|--|
| 7(c)(i) | <p>Treatment factor/variable – tyre brand A, B, C, D</p> <p>Blocking factor/variable – front or rear wheels</p> <p>or</p> <p>Table which could have rows/columns reversed</p> <table border="1"> <thead> <tr> <th></th> <th colspan="4">Brand</th> </tr> <tr> <th></th> <th>A</th> <th>B</th> <th>C</th> <th>D</th> </tr> </thead> <tbody> <tr> <td>Front Wheels</td> <td></td> <td></td> <td></td> <td></td> </tr> <tr> <td>Rear Wheels</td> <td></td> <td></td> <td></td> <td></td> </tr> </tbody> </table> <p>Measure the Tyre life for each factor combination</p> | | Brand | | | | | A | B | C | D | Front Wheels | | | | | Rear Wheels | | | | | B1 | 2.1a | Consideration of brands of tyre with front and rear wheels |
| | Brand | | | | | | | | | | | | | | | | | | | | | | | |
| | A | B | C | D | | | | | | | | | | | | | | | | | | | | |
| Front Wheels | | | | | | | | | | | | | | | | | | | | | | | | |
| Rear Wheels | | | | | | | | | | | | | | | | | | | | | | | | |
| 7(c)(ii) | 2-factor ANOVA | B1 | 2.1a | oe combine both Condone 2-way ANOVA | | | | | | | | | | | | | | | | | | | | |
| | Total | 17 | | | | | | | | | | | | | | | | | | | | | | |